

The Dirichlet process and clustering

Consider first a finite mixture model with K clusters, e.g., a GMM:

$$p(\pi | \alpha) = \text{Dir}(\pi; \frac{\alpha}{K} \vec{1}_K)$$

all 1 vector with K comp.

$$p(\theta^k | \lambda) = H(\theta^k; \lambda) \quad \text{conjugate prior to}$$

$$p(z_i = k | \pi) = \pi_k$$

$$p(x_i | z_i = k, \theta) = p(x_i | \theta^k)$$

We can write

$$\theta^k \sim H(\lambda)$$

$$x_i \sim F(\theta^{z_i})$$

Instead we can view this as

$$\theta_i \sim G(\theta) = \sum_{k=1}^K \pi_k \delta_{\theta^k}(\theta) \quad \begin{array}{l} \text{as above} \\ \text{Notice } \theta_i \text{ late } \theta^k \text{ cluster} \end{array}$$

$$x_i \sim F(\theta_i).$$

This gives s_k clusters. DP generalizes to "aug" k .

firs part

We define the DP using 2 concrete processes.

The stick-breaking process

Repeatedly:

- (1) break a stick, using a $\text{Beta}(1, \alpha)$ dist.
- (2) continue with non-picked part.

That is,

- start with length 1 stick
for $k=1 \dots \infty$ { - pick new fraction $\beta_k \sim \text{Beta}(1, \alpha) \propto (1-x)^{\alpha-1}$
} - let π_k be the absolute length

$$\pi_k = \beta_k \prod_{e=1}^{k-1} (1-\beta_e) = \beta_k \left(1 - \sum_{e=1}^{k-1} \pi_e\right)$$

left by earlier breaks

This is denoted

$$\pi \sim \text{GEM}(\alpha)$$

By adding

$$\theta^k \sim H(\lambda)$$

$$G(\theta) = \sum_{k=1}^{\infty} \pi_k \delta_{\theta^k}(\theta)$$

We get $G \sim DP(\alpha, H)$.

Note, by construction discrete.

We will tend to see the same θ^k many times, this induces the clusters.

Alternative: The Chinese restaurant process.

Let

$$p(\theta_{N+1} = \theta | \theta_{1:N}, \alpha, -) = \frac{1}{\alpha + N} (\alpha H(\theta) + \sum_{k=1}^K N_k f_{\theta^k}(\theta))$$

where

$K = \#$ different observations

$N_k = \#$ obs. of θ^k in $\theta_{1:N}$.

It is also possible to introduce a class variable z_i (so, $\theta_i = \theta^{z_i}$). Then

$$p(z_{N+1} = k | z_{1:N}, \alpha) = \begin{cases} \frac{\alpha}{\alpha + N} & \text{if } k \notin [K] \\ \frac{N_k}{\alpha + N} & \text{otherwise} \end{cases}$$

Interpretation

- $[K]$ tables/clusters
- new customer sits with others or at new table

- "large" tables have large prob. of growing.

Theo

$$P(\#\text{tables} \propto \log N) \xrightarrow{N \rightarrow \infty} 1.$$

DP mixture model

Up to now, we have generated clusters, or a distribution for each cluster.

Now also data from each sample cluster:

$$\pi \sim \text{GEM}(\alpha)$$

$$z_i \sim \pi \quad P(z_i=k|\pi) = \pi_k$$

$$\theta^k \sim H(\lambda)$$

$$x_i \sim F(\theta^{z_i}).$$

A collapsed Gibbs sampler for DPMM.

As before the states are of the form $z = (z_1, \dots, z_N)$.

Notice

$$p(z_i=k | z_{-i}, x, \alpha, \lambda) \propto p(z_i=k | z_{-i}, \alpha) p(x_i | x_{-i}, z_i=k, z_{-i}, \lambda)$$

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By interchangability (order of sampling irrelevant)

We can assume that i is the last customer.

Hence,

$$\begin{cases} \frac{N_{k,-i}}{\dots} & \text{if } k \in [K] \\ 0 & \text{otherwise} \end{cases}$$

$$p(z_i=k|z_{-i}, \alpha) = \begin{cases} \frac{\alpha}{\alpha+N-1} & \text{otherwise} \\ \frac{\alpha}{\alpha+N-1} & \dots \end{cases}$$

where

$K = \# \text{ cluster for } z_{-i}$

$N_{k,-i} = \# \text{ of } z_{-i} \text{ in cluster } k$.

(2)

Let $x_{n,-i} := \{x_j : z_j = k, j \neq i\}$.

Notice,

$$p(x_i|x_{-i}, z_{-i}, z_i=k, \lambda) = p(x_i|\bar{x}_{k,-i}, \lambda) = \frac{p(x_i, x_{n,-i}|\lambda)}{p(x_{n,-i}|\lambda)}$$

where den. and num. are marginal likelihood of data.

E.g.

$$p(x_i, x_{n,-i}|\lambda) = \int p(x_i|\theta) \prod_{j \neq i: z_j=k} p(x_j|\theta) H(\theta|\lambda) d\theta$$

If $z_i = K+1$ (a new cluster), then

$$p(x_i|x_{-i}, z_{-i}, z_i=k, \lambda) = \int p(x_i|\theta) H(\theta|\lambda) d\theta.$$

So for conj. prior H of F , we are typically ok.

Gaussian F with fixed variance and uniform H on its exp.

$$\int_0^1 \prod_i N(x_i; \mu, \sigma^2) d\mu$$

can be solved numerically.